

First microlensing candidates from the MEGA survey^{*} of M 31

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Abstract. We present the first M 31 candidate microlensing events from the Microlensing Exploration of the Galaxy and Andromeda (MEGA) survey. MEGA uses several telescopes to detect microlensing towards the nearby Andromeda galaxy, M 31, in order to establish whether massive compact objects are a significant contribution to the mass budget of the dark halo of M 31. The results presented here are based on observations with the Isaac Newton Telescope on La Palma, during the 1999/00 and 2000/01 observing seasons. In this data set, 14 variable sources consistent with microlensing have been detected, 12 of which are new and 2 have been reported previously by the POINT-AGAPE group. A preliminary analysis of the spatial and timescale distributions of the candidate events support their microlensing nature. We compare the spatial distributions of the candidate events and of long-period variable stars, assuming the chances of finding a long-period variable and a microlensing event are comparable. The spatial distribution of our candidate microlensing events is more far/near side asymmetric than expected from the detected long-period variable distribution. The current analysis is preliminary and the asymmetry not highly significant, but the spatial distribution of candidate microlenses is suggestive of the presence of a microlensing halo.

Key words. gravitational lensing – M 31: halo – dark matter

1. Introduction

One of the astrophysical solutions to the galactic dark matter problem would be the presence of a significant amount of undetected compact objects in the halos of galaxies. These MACHOs (Massive Astrophysical Compact Halo Objects) can be detected using gravitational microlensing (Paczynski 1986). According to gravitational lensing theory the measured brightness of a background source will temporarily increase if a massive compact object moves close enough through our line of sight towards the background source.

During the last decade, the MACHO (Alcock et al. 1993) and EROS (Aubourg et al. 1993) collaborations have been monitoring fields in the Large and Small Magellanic Clouds (LMC and SMC) in order to detect such events. After 5.7 years of observing the MACHO group found 13–17 microlensing

events towards the LMC, and concluded that up to 20% of the Milky Way dark halo may consist of compact objects of mass $0.15\text{--}0.9 M_{\odot}$ (Alcock et al. 2000). The EROS collaboration has found 3 microlensing events towards the LMC and puts strong constraints on the fraction of dark matter in the form of compact objects (Lasserre et al. 2000). The results of both groups are, however, consistent with $\sim 10\%$ of the dark halo mass consisting of compact objects of $\sim 0.5 M_{\odot}$ (e.g. Milsztajn 2002).

Looking for microlensing events in the nearby Andromeda galaxy (M 31) has several advantages over Magellanic Clouds searches (Crotts 1992; Baillon et al. 1993). Because of the geometry the microlensing optical depth can be up to ten times larger in parts of M 31. In combination with the extremely high density of background stars, this results in a highly enhanced microlensing rate. Due to the high inclination of the disk of M 31, it should be possible to use microlensing to constrain the mass contribution of compact objects to the dark halo (Baltz et al. 2003; Gyuk & Crotts 2000; Kerins et al. 2001). In the presence of a significant microlensing halo, the microlensing rate should be asymmetric, with more microlensing taking place towards the far side of the disk than to the near side,

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band in one night. The Julian date of the epoch is taken as the weighted average of the Julian dates of the individual frames.

- *Image subtraction*

From the single epoch images the high signal-to-noise reference image is subtracted, after photometric calibration and matching of the point spread function (PSF) between the images (Tomaney & Crotts 1996). The shape of the PSF is measured from bright, unsaturated stars in the images that are being matched. By dividing the PSFs in Fourier space a convolution kernel is calculated with which the better seeing image (usually the reference image) is degraded.

- *Variable object detection*

The resulting difference images are dominated by shot noise in which variable sources show up as positive or negative residuals, depending on the flux difference of the object between the single epoch image and the reference image. Due to fringing, the i' difference images are of poorer quality than the r' difference images. SExtractor (Bertin & Arnouts 1996) is used to detect residuals in all r' difference images from the first two observing seasons. The catalogs with residuals are cross-correlated to obtain a catalog with all variable objects in the surveyed fields. As a first selection to get rid of noisy detections, we demand that objects have to be detected in at least two epochs.

- *Lightcurves and epoch quality*

Lightcurves for the variable sources are obtained by performing PSF fitting photometry on the residuals in the difference images, using the PSF shape measured from the bright unsaturated stars. Several epochs turned out to give problematic difference images for a number of different reasons. Epochs with seeing worse than $2.0''$ do not give clean difference images; 9 epochs were discarded for this reason. In 7 epochs the images were overexposed, causing the PSF determination to fail, so these had to be discarded as well. During the 00/01 season there were problems with inaccurate guiding in a number of epochs. In 2 epochs the guiding failed completely and were useless.

Lightcurves were also produced at “empty” positions, i.e. positions where no variability was detected. Flat line fits were done to these empty lightcurves to check the error bars on the fluxes derived from statistics of the PSF fitting photometry. For each epoch, the distribution of the deviations from the flat lightcurve fits weighted by the error bar returned by the photometry routine was examined. In some cases this distribution showed broad non-Gaussian wings, and these epochs were discarded. Typically they were associated either with highly variable seeing between the individual exposures or inaccurate guiding. In other cases, the normalized error distribution was gaussian, but with dispersion higher than one. In these cases the error bars were renormalized appropriately.

After this, 27% of the 139 observed epochs in r' had been discarded and 37% of the 114 observed epochs in i' . The typical number of epochs that were left after the procedure described above are tabulated in Table 1 for each filter and field for the 99/00 and 00/01 seasons. From these epochs the lightcurves were constructed that were used for the analysis presented in this paper.

Table 1. Overview of the number of epochs used for seasons 1 and 2, field and filter. F1 is the north field and F2 is the south field.

	r'		i'	
	F1	F2	F1	F2
99/00	46	45	16	15
00/01	56	55	56	56
Total	102	100	72	71

2.3. Event selection

The final dataset consists of lightcurves of 118 424 variable sources, practically all of which are periodic variable stars. Finding candidate microlensing events in such a large number of lightcurves is no trivial problem. A procedure to select lightcurves that are compatible with microlensing must be aimed at recognizing the characteristics of a microlensing lightcurve while taking into account computing speed, and the quality of the available data. Since the quality of the i' band data was clearly poorer than the quality of the other bands, and since g' data were only available for the first season, we decided to use the r' data as the main basis for candidate selection. Another advantage of the r' data is that they have much better sampling density in the first season than do i' and g' data separately. Thus, the first steps of the filtering process involved only r' data, after which the i' data are used to further analyse the r' microlensing candidates.

A microlensing event caused by a single lens has a characteristic shape and a flat baseline. The selection steps based purely on the r' lightcurves are aimed at selecting lightcurves with flat baselines and well sampled, significant peaks which fit the characteristic microlensing shape well. A detailed description of these steps is given in Appendix A.1. To make sure these selection steps do a good job selecting microlensing events and rejecting long-period variable stars, Monte Carlo simulations of microlensing events and long-period variables were performed and used to fine tune the selection procedure. Based on just the r' data we select 1347 lightcurves.

Colour information can be used for the microlensing candidate selection for several reasons. Contrary to variable stars, microlensing events are intrinsically achromatic, meaning that the colour of the observed difference flux remains the same during the event. Furthermore, long-period variable stars that might be mistaken for microlensing are very red. Because of this, the brightness variations are also more pronounced in i' than in r' . Further selection steps, described in more detail in Appendix A.2, include a colour cut designed to reject long-period variables and a goodness of fit criterion for the i' lightcurve. For the latter, the i' lightcurves are fit with a standard microlensing lightcurve with the shape parameters found for the r' lightcurves, as a test for achromaticity.

This complete automated selection procedure results in 126 candidate microlensing events from the 99/00 and 00/01 observing seasons. Unfortunately, rather lenient χ^2 cuts had to be used for the goodness of fit criteria. This is due to the photometry being too sensitive to nearby variable sources, as described in more detail in Appendix A. Because of this, the sample still contains lightcurves with slightly variable baselines.

The lightcurves and the difference image residuals of these candidate events had to be inspected visually to establish whether any additional variability was intrinsic to the variable source or was caused by nearby sources. An effort to improve the photometry in the presence of neighbours is underway. The lightcurves with baseline variability that could not be attributed to nearby variable sources, in total 67, were thrown out. In a further 39 cases, the detected candidate event was actually caused by bad pixels and these were discarded as well, so after visual inspection, 21 events remain that are consistent with microlensing based on the 2 year lightcurves. Finally, part of the data from the third (01/02) season were used to extend the baselines of these 21 events, 7 of which showed variability in the third season. The final set of candidate microlensing events consists therefore of the 14 events described in the following section.

3. Results

The lightcurves of the 14 candidate microlensing events are shown in Fig. 2; their positions are plotted in Fig. 1. August 1st 1999 was taken as the zero point of the time scale in the lightcurve plots. The important fit parameters are tabulated in Table 2. For illustration we show in Fig. 4 thumbnails of the difference images and single epoch images for events 5, 7 and 11.

Inspection of the lightcurves in Fig. 2 indicates that the baselines are not always completely flat, this is especially the case for the i' lightcurves. When looking at the difference images, it becomes clear that for these events, secondary bumps in the lightcurve are caused by variable objects very close to the position of the event, rather than by variability in the events themselves. For example, the bump around day 100 in the lightcurve of candidate event number 2, is caused by a brightening source located only 4 pixels away from the event. In the i' data, this problem occurs more frequently because of the higher density of detected variable sources and the effect is stronger because of the stronger variability at longer wavelengths.

Not only can nearby variable sources result in trends in the lightcurves, they can also increase the scatter in the flux. The size of the aperture and the ring that is used for sky subtraction depends on the seeing, meaning that a nearby variable does not always influence the photometry and also not necessarily in the same way. As an example of this, we show in Fig. 3 the flux as a function of the seeing for the part of the baseline of event 10. Clearly, differences in seeing between the epochs will create extra scatter in the baseline.

Below we discuss the possible reasons for additional variability in each of the event lightcurves.

Event 1: this event has a bright periodic variable star at a distance of 5 pixels, which is probably the cause of the noisy baseline in both r' and i' .

Event 2: this event also has close variable neighbours. In the r' difference images a variable source is located at only 3 pixels from the event. This source peaks around day 100, causing a bump in the r' lightcurve. In the i' difference images there are two additional bright variables, both of them at about 10 pixels distance. These sources cause the noisy nature of the i' baseline.

Event 3: although the r' baseline is very clean, the i' baseline shows quite some variability, especially in the second season. Again this is caused by a very close variable source, at 3 pixels distance. This source has a bright episode between days 440 and 470, which causes the high fluxes in this period in the i' lightcurve.

Event 4: the baselines of this event look quite flat, but the peak in the i' data is not very symmetric. The sudden drop in flux after day 420 coincides with the sharp brightening of a bright variable source at 14 pixels distance.

Event 5: this event has several variable sources nearby. Unfortunately the closest one, at about 4 pixels, is also the brightest, which especially in the i' lightcurve causes a lot of noise. See Fig. 4 for thumbnails of this event.

Event 6: in the second season the baseline shows a periodic variability, which is caused by a periodic variable star at 6 pixels from the event. A second close variable, at 5 pixels, shows up only in the first season in r' , but pops up also in the second season in i' and is very bright in the first season. Together, these variables cause the i' baseline to be rather messy.

Event 7: the baselines of this very high S/N event are quite well behaved in both r' and i' , but there is a quite bright variable source located at 15 pixels, which may cause some deviations. However, the deviations in the i' data from the Paczynski fit between days 100 and 150 are too large to be caused by these near neighbours. Also from this event, some thumbnails are shown in Fig. 4.

Event 8: even though there is a variable source located at 7 pixels distance from this event, the baselines are quite well behaved.

Event 9: two faint variables are located at about 4 and 8 pixels distance from this event, and a bright one at 12 pixels. The lightcurves do not appear to be affected substantially.

Event 10: here the noisy baseline in i' and the two deviating points in the r' lightcurve around day 100 can be attributed to the same variable source 7 pixels away.

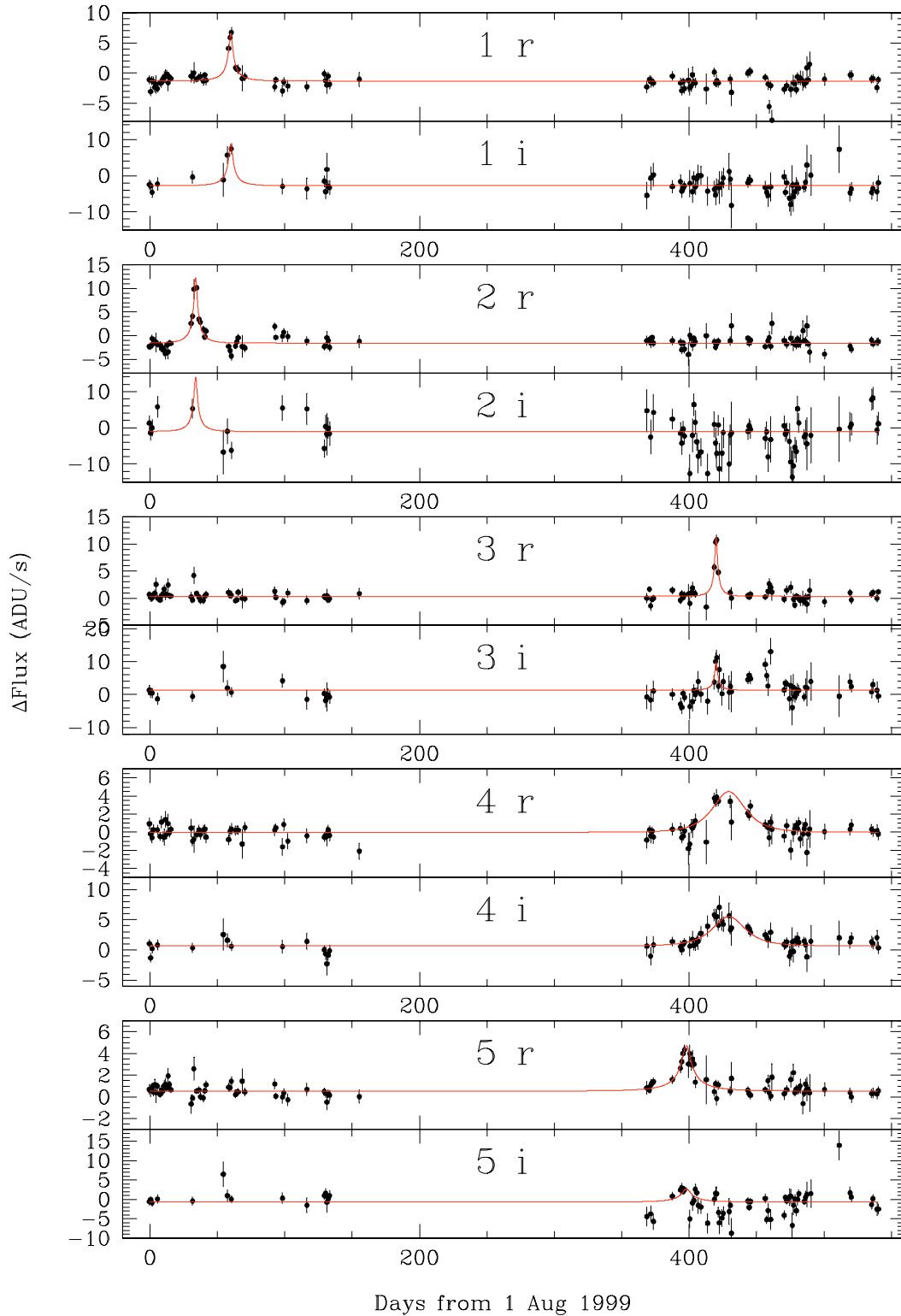
Event 11: while the r' baseline of this high amplification event is well behaved, the i' baseline is not completely flat. In the difference images there is a hint of a weak variable source a few pixels away, which is probably the cause of the small baseline wiggle. See Fig. 4 for thumbnails of this event.

Event 12: the r' lightcurve of this event is quite well behaved and there are no nearby variable sources. The reason for the noise in the i' lightcurve is not clear.

Event 13: a fairly bright variable source is situated 16 pixels away from this event, but the lightcurves in both r' and i' are well behaved.

Event 14: the noise in the lightcurves and more particularly the small bump in the i' lightcurve around day 370 is due to a variable source 3 pixels away from the event.

In Fig. 5 we plot the r' and i' fluxes for epochs in the peaks of the events for the cases where at least 3 epochs are available that have measurements in both bands. If the events are achromatic, as microlensing events should be, the points should lie on a straight line. Apart from events 5, 6 and 11, for all events the points follow a straight line and thus seem to be achromatic. This shows that the implicit achromaticity test we use during the candidate microlensing selection process by using



Days from 1 Aug 1999

Fig. 2. Lightcurves of the 14 microlensing candidate events. This page shows candidates 1 through 5, with the r' lightcurves plotted in the upper panel and the i' lightcurves in the lower panel. The Julian date of the zeropoint of the time axis is 2451393, which corresponds to August 1st 1999.

the parameters from the r' lightcurve fits for the i' fits, is working quite well.

Events 5, 6 and 11 are retained in the sample of candidate microlensing events. In the case of event 11 this is justified because the data are marginally consistent with a straight

line and there are only 3 points. The same is true for event 6, and furthermore this events was hardly detected at all in i' . For event 5 the i' photometry is heavily compromised by a nearby variable source, as can be clearly seen in Fig. 4.

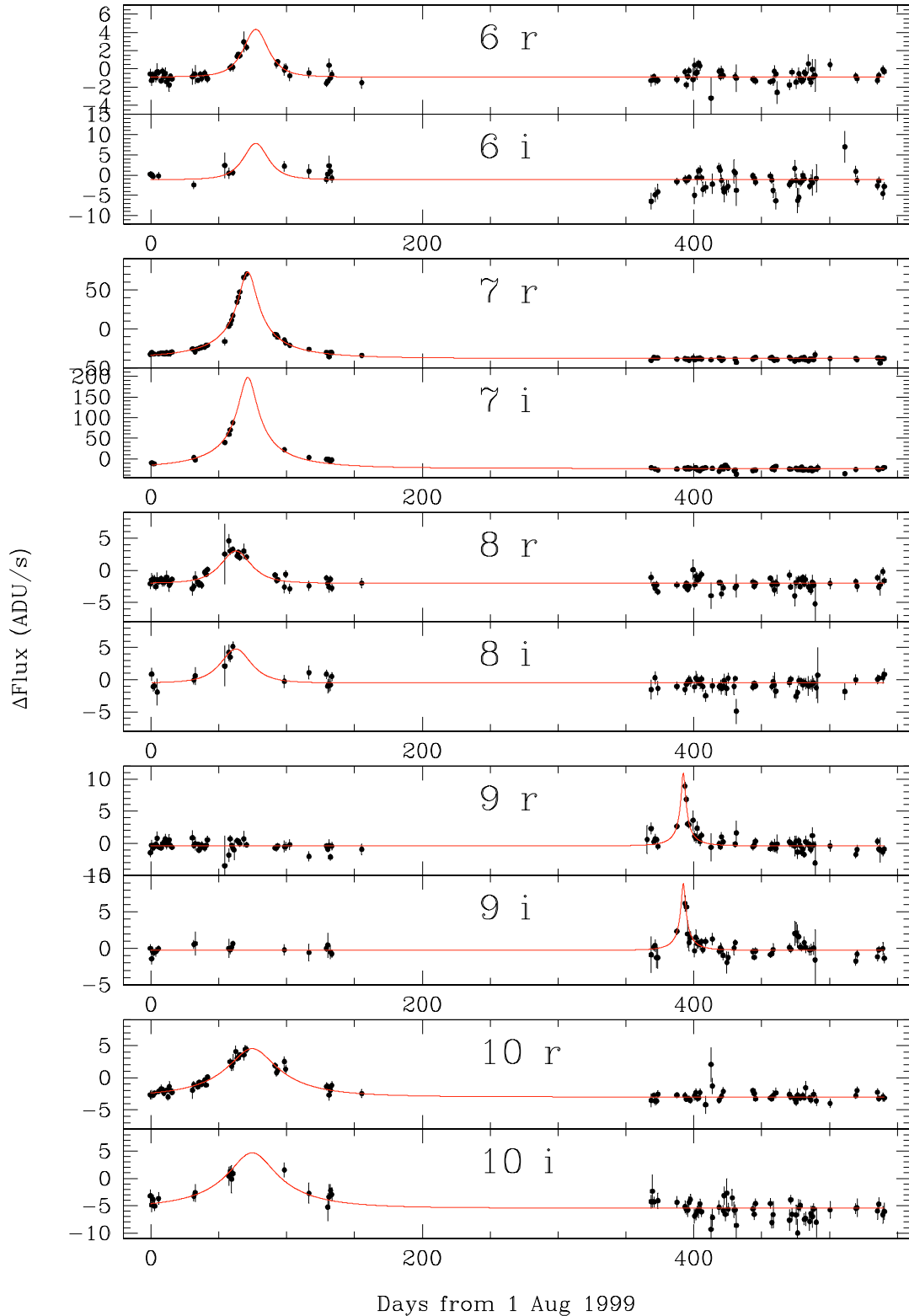


Fig. 2. continued. Candidates 6 through 10.

3.1. POINT-AGAPE events

The POINT-AGAPE group uses the same data for their microlensing survey towards M 31, although their pixel lensing technique is different (Paulin-Henriksson et al. 2003). Based on the same two-year dataset, they find four high signal-to-noise

microlensing events. Two of these events are also present in our sample, namely PA-99-N2, our event number 7, and PA-00-S4, our event number 11.

The fact that our sample of candidate microlensing events is much larger is caused by the severe selection criteria used by POINT-AGAPE, that limit the number of events found.

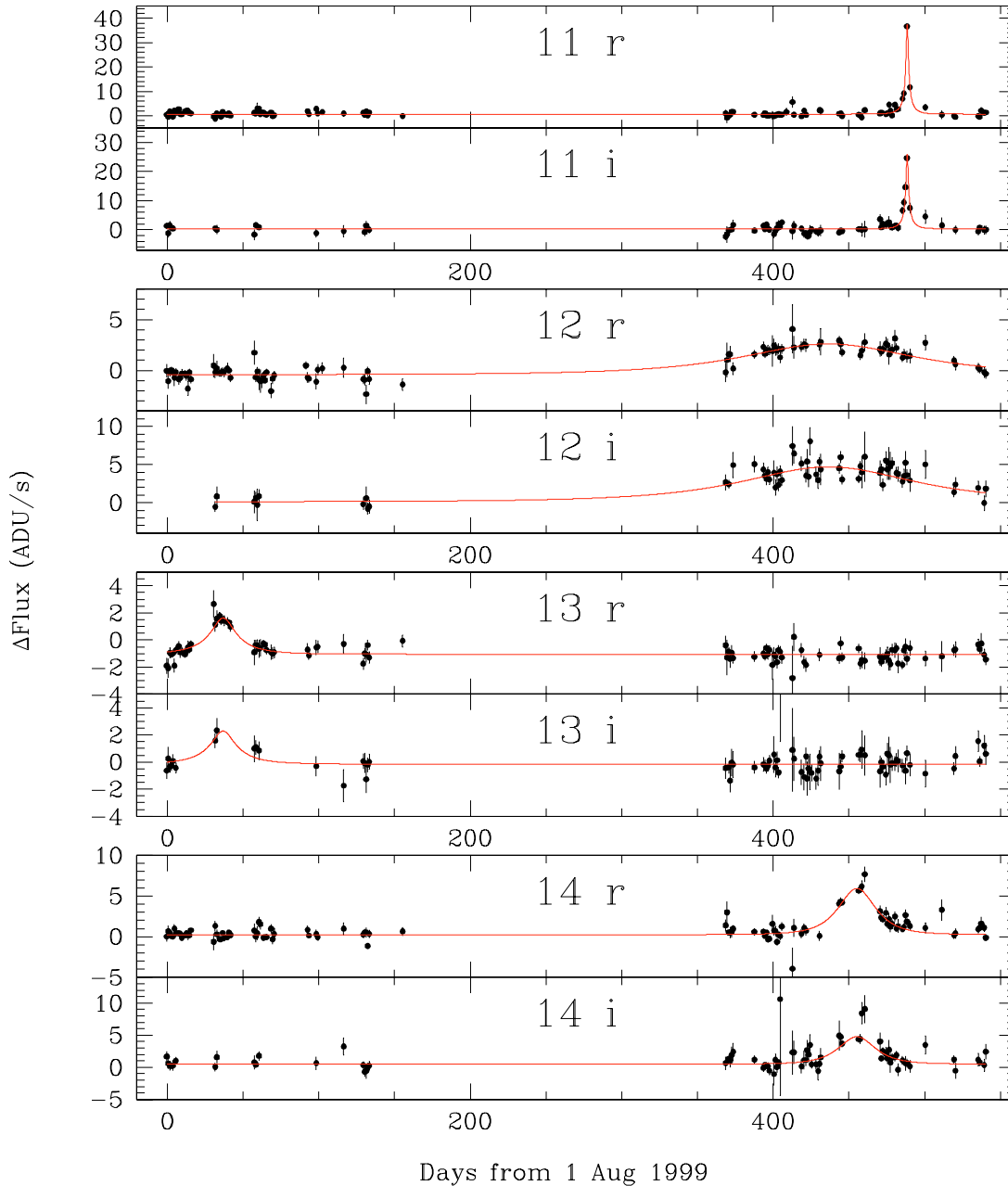


Fig. 2. continued. Candidates 11 through 14.

For example, POINT-AGAPE demands $FWHM$ timescales smaller than 25 days and much higher signal-to-noise than we. The following two events reported by the POINT-AGAPE team are not in our sample of candidate microlensing events.

- *PA-99-N1*

This event did not pass our filtering procedure because of too large variability in the i' baseline caused by a very closely variable star. Our lightcurves of this event are shown in Fig. 6.

- *PA-00-S3*

Very close to the bulge of M 31 the surface brightness becomes very high and our difference images become very noisy and of low quality. This event is located in a part of the field that is not used for the current analysis for this reason. In Fig. 1 the region that is not used in our analysis is indicated.

4. Discussion

Due to the nature of M 31 microlensing, it is very difficult to determine the nature of the lensing object in individual microlensing events. In classical microlensing the mass of the lens, the relative motion of the lens with respect to the source and the distance to the lens are usually unknown and degenerate parameters. In difference-image lensing there is an additional unknown, namely the unlensed flux of the source. If a conclusion about the existence and/or nature of a MACHO population in the halo of M 31 is to be drawn from a microlensing survey, statistical methods have to be used.

By modeling the population of stars that can be lensed and the population of objects that can act as lenses, predictions can be made about the rate of occurrence, the spatial distributions,

Table 2. Fit parameters for the 14 candidate microlensing events. The t_{MAX} and t_{FWHM} values are based on full parameter fits to the r' lightcurves. Also listed are the magnitude of the difference flux at maximum amplification and the $r'-i'$ colours where possible. For the short events, i.e. t_{FWHM} in the order of a few days, the uncertainties are large, because the time sampling of the data is relatively sparse compared to the width of the peak. Peak times are in days after August 1st 1999.

Candidate event	RA (J2000)	Dec (J2000)	t_{MAX} (days)	t_{FWHM} (days)	$\Delta r'$ (mag)	$r'-i'$ (mag)
MEGA-ML 1	0:43:10.54	41:17:47.8	60.1 ± 0.1	4.2 ± 4.3	22.2 ± 1.1	1.1 ± 1.5
MEGA-ML 2	0:43:11.95	41:17:43.6	34.08 ± 0.08	4.6 ± 0.6	21.6 ± 0.3	
MEGA-ML 3	0:43:15.76	41:20:52.2	420.06 ± 0.05	2.6 ± 2.2	21.8 ± 1.2	0.4 ± 1.5
MEGA-ML 4	0:43:04.08	41:26:15.6	429.7 ± 0.2	29.1 ± 1.0	22.8 ± 0.2	0.8 ± 0.3
MEGA-ML 5	0:44:48.95	41:22:59.3	398.1 ± 0.2	9.4 ± 4.1	22.9 ± 0.8	0.4 ± 1.0
MEGA-ML 6	0:44:50.97	41:24:42.4	77.3 ± 0.2	22.9 ± 0.7	22.6 ± 0.2	
MEGA-ML 7	0:44:20.89	41:28:44.6	71.20 ± 0.06	21.6 ± 0.7	19.3 ± 0.2	1.4 ± 0.2
MEGA-ML 8	0:43:24.53	41:37:50.4	63.1 ± 0.2	27.4 ± 0.9	22.7 ± 0.2	0.8 ± 0.2
MEGA-ML 9	0:44:46.80	41:41:06.7	392.3 ± 0.2	3.8 ± 1.6	21.8 ± 0.8	0.4 ± 1.0
MEGA-ML 10	0:43:54.87	41:10:33.3	74.7 ± 0.4	46.8 ± 4.4	22.2 ± 0.3	1.0 ± 0.3
MEGA-ML 11	0:42:29.90	40:53:45.6	488.6 ± 0.05	2.0 ± 0.3	20.5 ± 0.2	0.3 ± 0.2
MEGA-ML 12	0:41:26.90	40:49:42.1	436.6 ± 0.6	131.0 ± 9.4	23.2 ± 0.3	1.0 ± 0.3
MEGA-ML 13	0:43:02.49	40:45:09.2	37.3 ± 0.5	22.8 ± 3.8	23.3 ± 0.3	0.5 ± 0.4
MEGA-ML 14	0:43:42.53	40:42:33.9	455.1 ± 0.3	28.1 ± 1.4	22.5 ± 0.2	0.3 ± 0.4

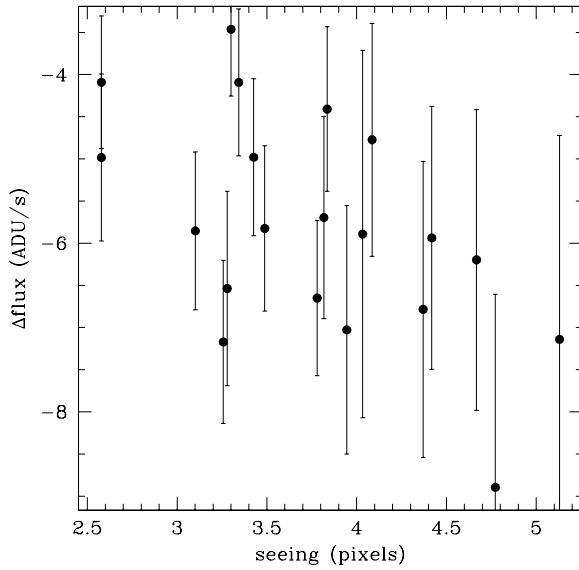


Fig. 3. Data points from the baseline of event 10, but now the $\Delta Flux$ is plotted as function of the seeing. During the period from which these points are taken, a variable source located about 7 pixels away from the event caused a strong residual in the difference images. There clearly is a trend with worse seeing corresponding to lower fluxes.

and the timescales of microlensing events. All these data can be used to constrain the lens populations responsible for observed microlensing events (Baltz et al. 2003, hereafter BGC; Kerins et al. 2001).

To compare our sample of candidate microlensing events with rate maps as published by BGC, the detection efficiencies and observing characteristics have to be carefully modeled, which will be done in a future paper. Instead of a full statistical analysis of the current sample, we restrict our discussion to a simple analysis of the spatial and timescale distributions of the

candidate events to see if they are consistent with predictions for halo lensing and lensing by stars in M 31.

4.1. Timescale distribution

The timescale of a microlensing event depends on the relative velocities of the source and the lens and the size of the Einstein radius, which in turn depends on the mass of the lens and the geometry (Eq. (A.2)). In the case of M 31 microlensing we can assume that $D_{\text{OL}} = D_{\text{OS}}$ and since the Einstein ring crossing time $t_{\text{E}} = 2R_{\text{E}}/v_{\perp}$ we get:

$$t_{\text{E}} \propto M_{\text{L}}^{1/2} D_{\text{LS}}^{1/2} v_{\perp}^{-1} \quad (1)$$

where v_{perp} is the relative speed between lens and source projected on the sky. The t_{FWHM} that we measure is related to t_{E} according to Eq. (A.5).

For self-lensing, i.e. lensing by stars, this translates into a timescale distribution that varies spatially. In the bulge, velocities are relatively high and distances between lens and source relatively short (compared to lensing by halo objects), therefore event durations will be relatively short. In the disk, relative velocities are much slower and event durations longer than in the bulge. On top of that, the timescales are expected to rise when going to larger radii, since the random velocities of the stars in a galactic disk are proportional to the root of the surface density (Bottema 1993). Roughly speaking we can say that average self-lensing timescales are inversely correlated with surface brightness; in higher surface brightness regions the timescales are expected to be shorter.

For halo lensing, relative velocities are higher than for disk-disk lensing, but because of the larger lens-source distance the event durations are still relatively long. Because the relative velocities and distances do not change much within the observed field, the timescale distribution of halo-lensing events should

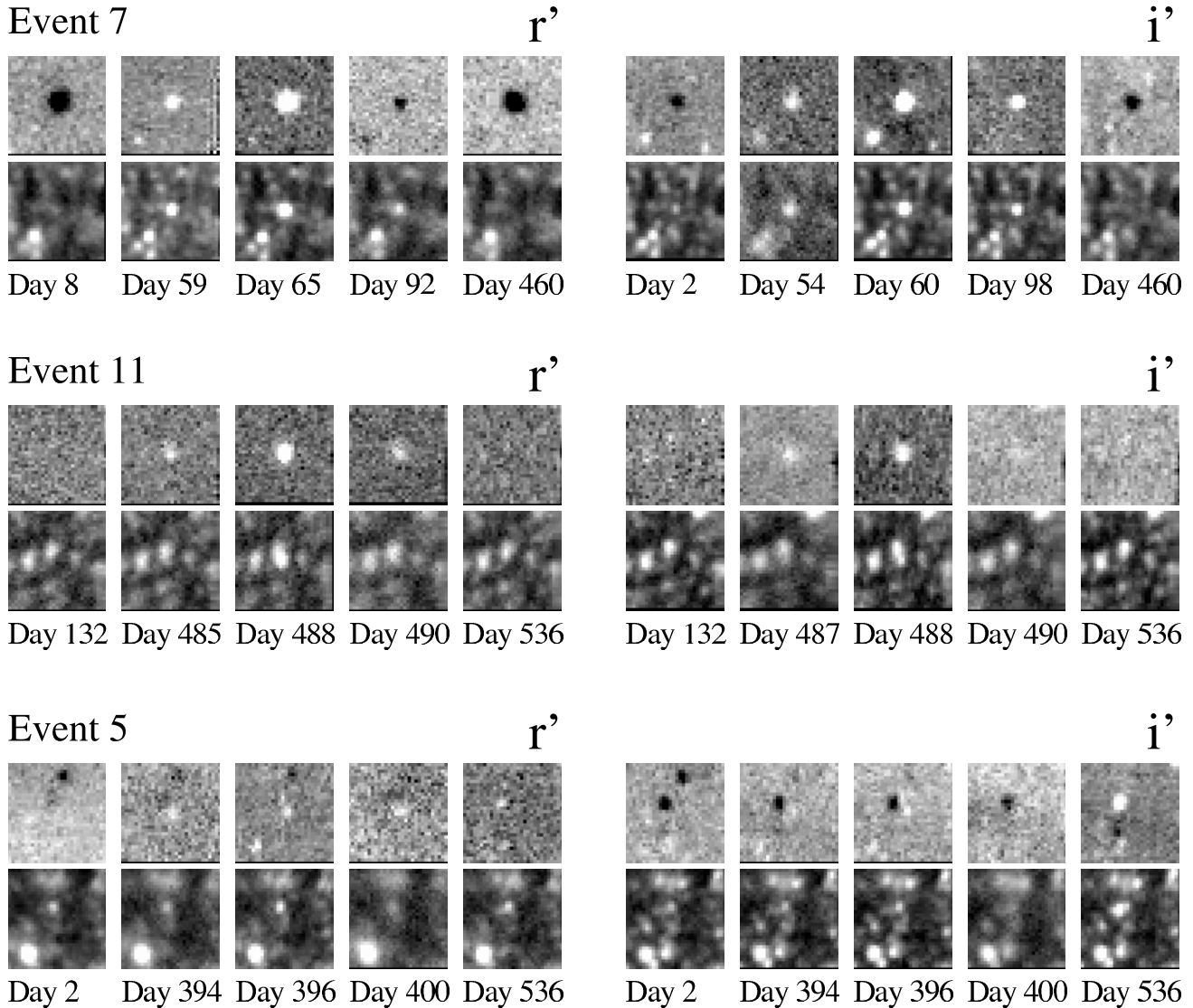


Fig. 4. Thumbnails of three events, chosen to span a wide range of peak fluxes. The uppermost panel shows event 7, by far the brightest event, going down in flux through event 11 (*middle panel*) to the low signal-to-noise event 5 (*lower panel*). For each event we show thumbnails of the difference images and single epoch images in both r' and in i' . Thumbnails were chosen from three epochs during the microlensing event, i.e. in the peak of the lightcurve, and from one epoch before and one epoch after the event. Event 7 (*upper panel*) stands out very clearly in both the difference images and the single epoch images in both bands. In the i' thumbnails a variable star is clearly visible in the lower left corner. The event in the middle panel, event 11, has a peak flux only one third of the peak flux of event 7, but still stands out clearly in the difference images. In the single epoch images, the event is only visible during maximum amplification. Event 5, shown in the lower panel is an example of the fainter events, with a peak flux almost an order of magnitude smaller than event 11. As can also be seen from the lightcurves (Fig. 2), the event is much brighter in r' than in i' . In the r' difference images the residual is more obvious than in the i' difference images. The nearby variable source that influences the i' photometry is clearly visible just above and to the left of the event. In the single epoch images, the event is hardly visible at all without difference image techniques.

be spatially uniform. Plots of the expected timescale distributions for different lines of sight towards M 31 are given by BGC for stellar self-lensing and halo lensing combined and reflect these trends.

In Fig. 7 we plot for our 14 microlensing candidates the logarithm of t_{FWHM} versus the r' surface brightness at their locations. The candidate events close to the center of M 31, in locations with surface brightness brighter than $19.5 \text{ mag}_{r'} \text{ arcsec}^{-2}$, all have event durations in the order of a few days. Most of the candidates in the low surface brightness regions have significantly longer timescales. That this distribution is exactly what

is expected for microlensing and the fact that the timescales are consistent with the timescales predicted by BGC are further indications that our sample of candidates consists of actual microlensing events.

4.2. Spatial distribution

The expected spatial distribution of stellar self-lensing events is strongly concentrated to the center because both the density of possible source stars and lenses is highest there. Furthermore,

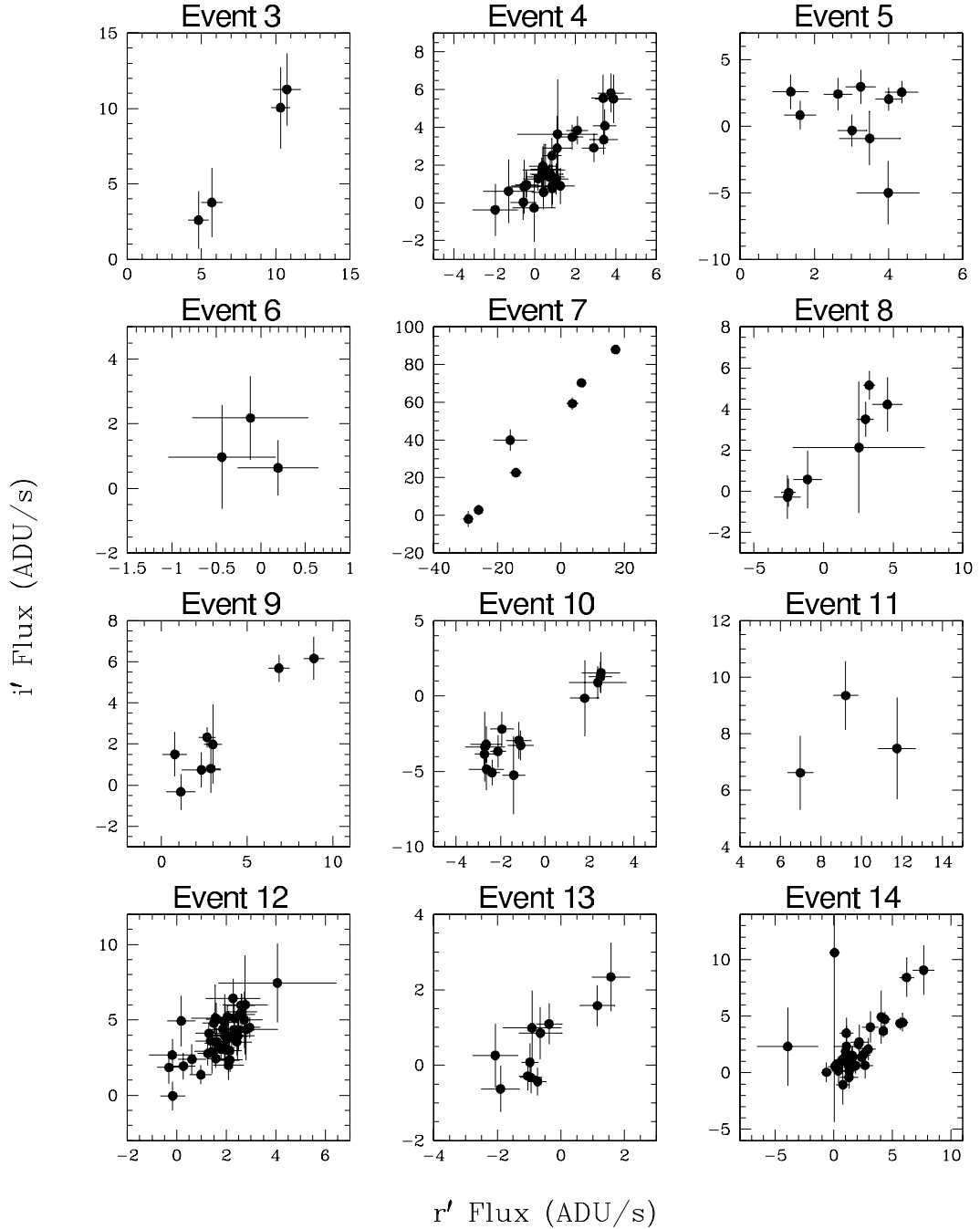


Fig. 5. Plots of the r' -band versus the i' -band flux for epochs located in the peaks of the candidate microlensing events. Epochs are plotted that lie within 2 times the t_{FWHM} from the fitted peak time. Only events are plotted that have at least 3 epochs in the peak with datapoints in both bands. The points from events 3, 4, 7, 8, 9, 10, 12, 13 and 14 lie on a straight line and thus are consistent with achromaticity. Events 6 and 11 are marginally consistent with achromaticity. Event 5 looks clearly problematic, but the i' -band photometry is strongly influenced by a very close-by variable star.

the spatial distribution is nearly symmetric from the far to the near side of the disk of M 31 (BGC), so in the case of only self-lensing the number density of events should be the same at the same radius on both the far and near side. The stellar halo of M 31 has not been taken into account by BGC, but is not expected to change the symmetry significantly, since, like the bulge, it acts both as a lens and as a source population. If the spheroid has the same M/L as the disk, for every event caused by a spheroid-lens and disk-source pair, there is a disk-lens and

spheroid-source pair event, as is the case with disk-bulge and bulge-disk lensing (BGC). In case of heavy extinction in the disk, the spheroid could induce an asymmetry in the lensing rate.

Dark halo lensing on the other hand will induce an asymmetric microlensing distribution, because it only acts as a lens population. Due to the longer line-of-sight through the M 31 halo towards the far side of the disk, the microlensing optical depth is much higher on the far side than on the near side.

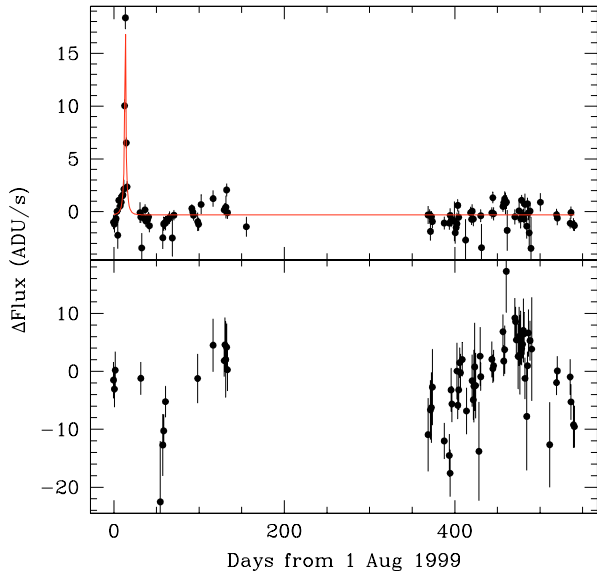


Fig. 6. The lightcurves in r' (upper panel) and i' (lower panel) of the microlensing event PA-99-N1, reported by POINT-AGAPE. In the i' data the variability in the baseline, caused by nearby variables, is high enough to give a χ^2 which is too high to pass our selection procedure.

Therefore, an asymmetry in the density of microlensing events between the far and near side is a strong indication of a microlensing halo. Halo lensing is also more likely near the center because of the higher surface density of possible source stars, but drops off much more slowly, since the surface density of halo lenses should not drop off very quickly.

Without knowing the detection efficiency of microlensing events, the number of detected candidate events is hard to interpret quantitatively. Since the detection efficiency might vary with position, their spatial distribution is also hard to interpret. However, if we assume that the spatial distribution of variable stars is symmetric over M 31, we can get an idea of the variation of the detection efficiency with position by looking at the spatial distribution of the detected variable stars. The variable stars that are most likely to be mistaken for microlensing are variables with periods longer than 150 days. In Table 3 we present the detected numbers of variables with periods between 150 and 600 days, for $300 \times 324''$ subregions in our field of view. For each subregion the actual area, corrected for area lost due to bright stars and diffraction spikes, is also tabulated. In Fig. 8 we show the corresponding variable star density map, with the positions of the 14 candidate events plotted on top.

To select the long-period variables (LPVs) from the complete set of lightcurves, all lightcurves were checked for periodic behaviour using the Numerical Recipes (Press et al. 1992) algorithm based on the Lomb method (1976) for spectral analysis of unevenly sampled data. The lightcurves for which a periodicity between 150 and 600 days was found with very high significance were selected. In total 32 841 variables were selected, excluding the part of the south field close to the bulge and one other small part at the edge of the south field, where a very bright foreground star is located.

In the spatial distribution of the LPVs we see no surprises. The surface density of these objects is correlated with the

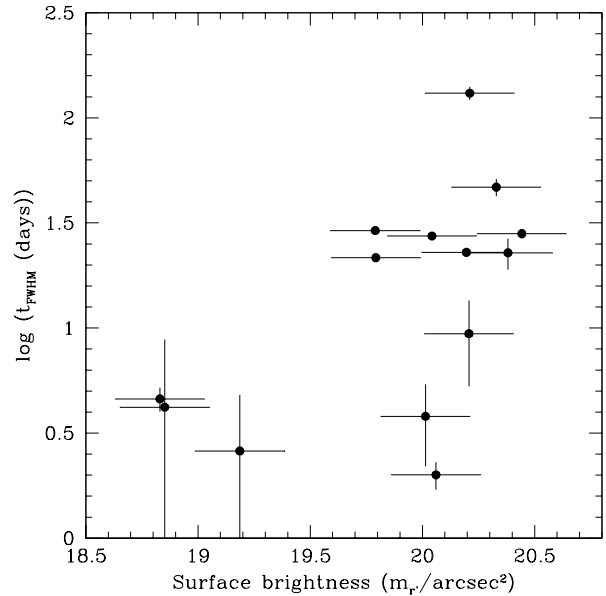


Fig. 7. For the 14 candidate microlensing events, the logarithms of the $FWHM$ timescales are plotted against the local r' surface brightnesses. At high surface brightnesses the timescales tend to be shorter than in low surface brightness regions.

surface brightness. Since we are interested in a possible asymmetry in the distribution of the candidate microlensing events, we compare the relative distributions of variables and events over the near and the far side of M 31. We assume that the LPVs are a good tracer of the population of possible microlensing source stars and that the chances of detecting LPVs and microlensing events are the same. The detection efficiencies of both will depend on the surface brightness of the disk and be subject to extinction effects.

To investigate the different parts of the far and near side of the M 31 disk that we sample, we divide each subregion in Table 3 in 4 smaller $150 \times 162''$ pieces. In the upper panel of Fig. 9 we plot the distribution of the LPV surface densities of these small subregions. The smaller, black histogram shows the distribution of the far side candidate events over the LPV density bins. If the microlensing rates are intrinsically near/far side symmetric, as we assume the LPVs to be, the expected number of candidate events on the near side can be determined for each LPV density bin, by scaling the number of far side events. We can only do this for the low surface density bins, since the high surface density bins are not populated sufficiently. However, since close to the center of M 31 the microlensing optical depth is dominated by stellar self-lensing they should not affect the symmetry of the spatial distribution. The lower panel of Fig. 9 shows the distribution of the subregions on the near side over the LPV density bins. The shaded histogram shows the distribution of the far side candidate events scaled to the near side LPV density distribution, and the black histogram the actual near side candidate events.

Whereas 6 near side candidate events are predicted from the distribution of the 9 far side events in areas with fewer than 10 LPVs per square arcsec, only 2 are detected. Thus, the distribution of candidate events seems to be more near/far

Table 3. Number counts of detected long-period variable (LPV) stars. For each subregion, the coordinates of the center, the number of LPV's and the area, corrected for area lost to bright stars and spikes, are given.

RA (J2000)	Dec (J2000)	LPV #	Area (\square')	RA (J2000)	Dec (J2000)	LPV #	Area (\square')
0:44:43.45	41:24:49.3	444	26.00	0:42:13.67	41:36:13.5	241	26.18
0:44:42.78	41:19:51.5	390	25.91	0:42:40.14	41:36:14.6	269	26.25
0:44:14.82	41:24:55.5	577	26.37	0:42:13.68	41:30:50.9	248	26.10
0:44:14.21	41:19:57.3	534	26.27	0:42:40.13	41:30:51.2	349	26.08
0:43:46.07	41:25:1.6	666	25.92	0:42:13.77	41:25:28.0	401	26.02
0:43:45.53	41:20:3.3	754	26.38	0:42:40.19	41:25:27.6	739	26.25
0:43:17.27	41:25:7.9	887	26.16	0:42:13.92	41:20:6.0	761	25.92
0:43:16.80	41:20:9.4	1340	26.36	0:42:40.32	41:20:5.0	1428	24.34
0:44:44.58	41:47:58.6	369	26.47	0:44:44.39	41:36:30.8	422	26.21
0:44:44.15	41:43:4.1	446	26.42	0:44:43.88	41:31:32.6	463	25.64
0:44:15.93	41:48:8.2	253	26.26	0:44:15.69	41:36:38.1	317	25.57
0:44:15.46	41:43:13.1	378	26.51	0:44:15.20	41:31:39.4	436	26.14
0:43:47.16	41:48:16.3	202	26.40	0:43:46.87	41:36:44.7	412	25.60
0:43:46.65	41:43:20.9	244	25.38	0:43:46.39	41:31:45.6	519	26.03
0:43:18.32	41:48:23.1	239	26.64	0:43:18.00	41:36:50.4	351	26.15
0:43:17.78	41:43:27.5	186	25.53	0:43:17.53	41:31:51.3	599	25.97
0:43:52.75	40:48:54.8	139	26.78	0:41:24.61	41:0:17.3	383	26.24
0:43:52.20	40:43:57.2	100	26.61	0:41:50.85	41:0:18.2	480	25.75
0:43:24.36	40:48:59.6	203	26.52	0:41:24.63	40:54:54.3	375	26.35
0:43:23.88	40:44:1.4	155	26.54	0:41:50.83	40:54:54.9	390	25.41
0:42:55.86	40:49:4.3	362	26.57	0:41:24.71	40:49:31.4	357	26.36
0:42:55.46	40:44:5.9	246	26.58	0:41:50.86	40:49:31.5	402	25.67
0:42:27.33	40:49:8.9	421	26.65	0:41:24.88	40:44:9.4	393	26.53
0:42:26.99	40:44:10.5	263	26.53	0:41:50.97	40:44:8.8	425	25.41
0:43:52.75	40:48:54.8	550	26.01	0:43:53.25	41:0:37.5	284	26.30
0:43:52.20	40:43:57.2	492	26.42	0:43:52.89	40:55:39.1	202	26.07
0:43:24.36	40:48:59.6	877	26.05	0:43:24.79	41:0:43.1	406	25.57
0:43:23.88	40:44:1.4	594	25.80	0:43:24.46	40:55:44.3	277	26.22
0:42:55.76	40:47:49.6	690	13.16	0:42:56.21	41:0:47.7	540	25.43
0:42:55.46	40:44:5.9	799	25.49	0:42:55.89	40:55:48.7	474	26.33
0:42:26.99	40:44:10.5	847	25.06	0:42:27.61	41:0:51.6	568	26.45
				0:42:27.29	40:55:52.4	564	25.73

asymmetric than the distribution of variable stars, which is an indication that it is a separate population of objects and not some kind of leakage from the population of LPVs. As argued before, the most likely explanation of such an asymmetry of microlensing events is dark halo lensing. The number of candidate events is still quite small, but the chance of finding a far:near split as strong as 9:2 is only 12%. It has been argued by Paulin-Henriksson et al. (2002) that candidate event 11 is possibly caused by a stellar lens in M 32 and not by a lens in the halo. Doing the same calculation without candidate event 11, the chance of finding the 8:2 ratio is 16%.

5. Conclusions

In this paper we present the first 14 microlensing candidates from the MEGA survey and a preliminary discussion of their properties. Our photometry turns out to be sensitive to nearby variable sources in a way that causes us to miss out on events

like PA-99-N1. It also leads to the necessity of a by-eye selection step in the filtering of the lightcurve database. Before the photometry is improved, enabling the filtering procedure to be automated and a detailed efficiency calculation to be done, we cannot draw firm conclusions based on the spatial distribution of our sample of candidate microlensing events. However, the spatial distribution of our candidate microlensing events is more near:far asymmetric, than would be expected from the spatial distribution of the detected long-period variable stars. Although the significance of this asymmetry is low, due to the small number of events and the lack of a detailed detection efficiency, the spatial distribution of the candidate microlensing events is suggestive of the presence of a microlensing halo.

Thus far we have only extracted candidate microlensing events from the 1999/00 and 2000/01 seasons. As a next step the remaining seasons will also be searched for microlensing events. Since we also observe with other telescopes than the INT, much more events will likely be detected when the data

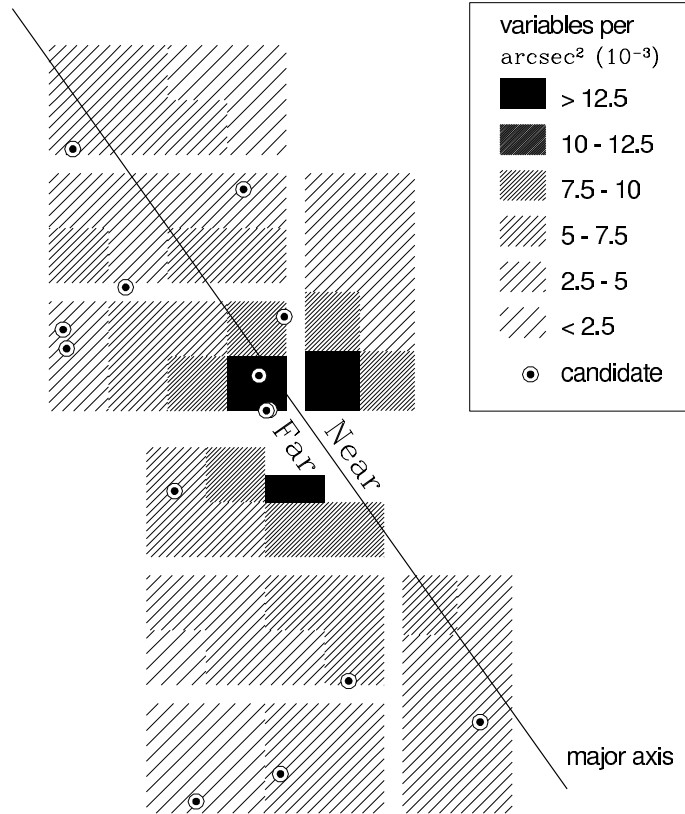


Fig. 8. Surface density map of long period variable stars in the surveyed area. Each chip was subdivided in 2×4 rectangles ($300 \times 324''$), in which the number of variables with periods longer between 150 and 600 days was counted. Only stars with accurately determined periods were used and the edges of the chips were avoided to ensure only well sampled lightcurves were used, leaving 32 841 variables. The positions of the 14 microlensing candidates are indicated, as well as the major axis of M 31.

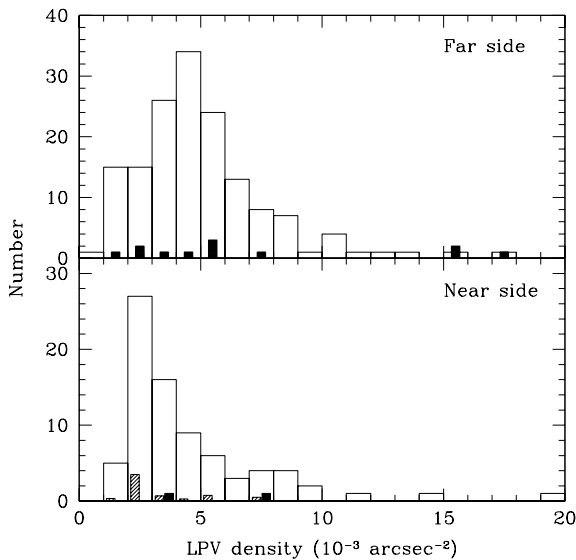


Fig. 9. *Upper panel:* histogram of the surface densities of long-period variable stars (LPVs) in small $150 \times 162''$ regions on the far side of the M 31 disk. The thin, black bars show in which LPV surface density bins the far side microlensing candidate events are located. *Lower panel:* histogram of the surface densities of LPVs in small $150 \times 162''$ regions on the near side. The thin, shaded bars show the numbers of candidate microlensing events predicted by scaling the numbers of far side events with the ratio of regions on the near and far side in each LPV density bin. In total 6 events are expected on the near side, but only 2 are detected, shown by the thin, black bars.

from the different telescopes are combined. Especially the sensitivity to short timescale events will benefit from merging the INT data with the data obtained with the telescopes of the MDM observatory, as they are supplementary in time coverage. Fainter events will also become easier to detect because the MDM data go deeper than the INT data, as do the data obtained with the KPNO 4m telescope.

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Appendix A: Microlensing event selection

A microlensing event caused by a single lens has a characteristic shape and a flat baseline. Such a standard microlensing lightcurve is called a Paczynski lightcurve (Paczynski 1986) and is described by

$$F(t) = F_0 \times A(t) = F_0 \times \frac{u^2 + 2}{u \sqrt{u^2 + 4}} \quad (\text{A.1})$$

where F_0 is the baseline, unlensed flux, $A(t)$ is the amplification, and u is the projected distance between the lens and the

source, in units of the Einstein radius. This Einstein radius depends on the geometry of the system and the mass of the lens and in the lens plane is given by:

$$R_E = \sqrt{\frac{4Gm D_{OL}D_{LS}}{c^2 D_{OS}}} \quad (\text{A.2})$$

where m is the lens mass and the D s are the distances between observer, lens and source. Note that the amplification is always 1 or higher and independent of wavelength, meaning that microlensing is in principle achromatic and conserves the colour of the source. If the relative motion of lens and source is taken to be a uniform motion, then u can also be written as:

$$u(t) = \sqrt{u_{\text{MIN}}^2 + \left(\frac{t - t_{\text{MAX}}}{t_E}\right)^2} \quad (\text{A.3})$$

where u_{MIN} is the impact parameter, t_{MAX} the time of maximum amplification and t_E the Einstein time. This is defined as the time it would take the source to cross the Einstein radius. Since we are measuring only the flux difference of variable objects with respect to a template image, Eq. (A.1) transforms into:

$$\Delta F(t) = \Delta F_{\text{BL}} + F_0 \times (A(t) - 1) \quad (\text{A.4})$$

where ΔF_{BL} is the baseline flux minus the flux on the reference image. As the crowding in M 31 prevents the unlensed source flux from being measured, ΔF_{BL} is an unknown parameter that has to be fitted.

Furthermore, since the baseline flux is unknown and the microlensing event resolved only while magnified, it is very difficult to measure t_E and u_{MIN} (Gould 1996; Baltz & Silk 2000), consequently these two parameters are highly degenerate, as shown in Fig. A.1. Instead, we fit the width of the peak at half of the maximum flux t_{FWHM} (Gondolo 1999), which is related to t_E by:

$$t_{\text{FWHM}} = t_E \times w(u_{\text{MIN}}) \quad (\text{A.5})$$

$$\text{where } w(u_{\text{MIN}}) = 2 \sqrt{2f(f(u_{\text{MIN}}^2)) - u_{\text{MIN}}^2} \quad (\text{A.6})$$

$$\text{with } f(x) = \frac{x+2}{\sqrt{x(x+4)}} - 1. \quad (\text{A.7})$$

We perform a four-parameter fit by letting t_{FWHM} replace t_E and u_{MIN} . To save computer time we assume a fixed value for u_{MIN} and let t_E float to fit for t_{FWHM} . Since there is some difference in the shape of the peak depending on u_{MIN} , we make fits for three values of u_{MIN} , namely 0.1, 0.3 and 1.0. These fits are sufficient for our selection procedure. Figure A.1 shows that this range covers the range of possible shapes very well.

A.1. Selection of r' lightcurves

The selection of microlensing candidates was based on the goodness of fit of the standard Paczynski lightcurve fits to the 99/00 and 00/01 data. Based on the parameters of the best fits, the lightcurves are filtered in several steps, described below.

• Peak sampling

To be able to constrain event parameters it is important to sample the peak of the lightcurve well. We demand that at least the

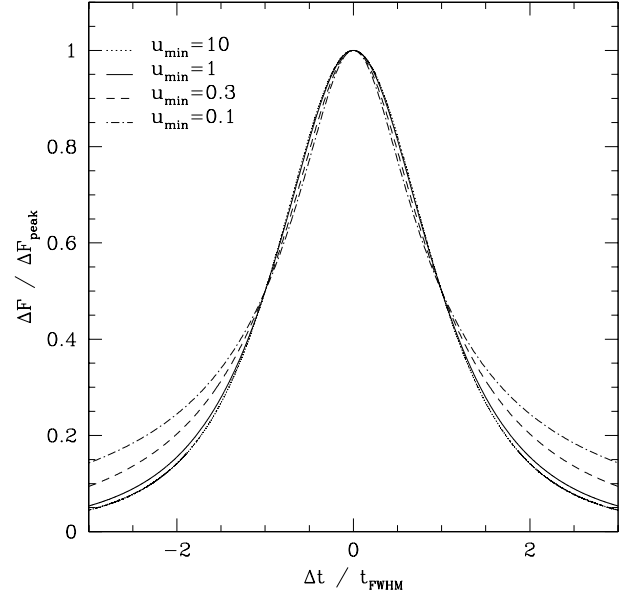


Fig. A.1. Differences in the Paczynski lightcurves for different values of the impact parameter u_{MIN} . The peak shape does not depend strongly on u_{MIN} , except from in the wings. In M 31, where microlensing events are often only resolved while significantly magnified, the wings are usually not strongly constrained, meaning that it is difficult to measure u_{MIN} .

part of the peak where the flux is more than half of the peak flux lies completely within an observing season. For this the t_{MAX} and t_{FWHM} from the Paczynski fits are used. We also demand that there are at least 2 data points above 25% of the fitted peak flux. Furthermore, we calculate a “peak weight” P defined as the sum of all data points within the peak:

$$P = \sum_{i \in \text{peak}} \frac{\Delta F_i - \Delta F_{\text{BL}}}{\sigma_i} \quad (\text{A.8})$$

where ΔF_i is the difference flux, σ_i the error on the difference flux, and ΔF_{BL} is the difference flux of the baseline, as given by the Paczynski fit. All the points that are within $2 t_{\text{FWHM}}$ of t_{MAX} are considered to be in the peak and are used for this calculation. All points farther away from the center of the peak are considered to be part of the baseline. A minimum “peak weight” P of 25 is demanded.

• Peak significance

To exclude spurious detections, a minimum peak significance is demanded. The χ^2 value of a flat line fit to the lightcurve is compared to the χ^2 of the Paczynski fit. The difference between these two values has to be larger than 100 for selection.

• Peak width

Both a minimum and a maximum peak width are used. The minimum t_{FWHM} of 1 day serves as an extra filter against spurious detections. Events with FWHM timescales longer than 150 days are also excluded, since with the current dataset the baseline is not well sampled. In practice, fits that give these very long timescales are usually caused by lightcurves that continue to rise at the end of an observing season.

• Baseline flatness

Microlensing events have flat baselines, contrary to periodic

variables. The flatness of the baseline is checked by the goodness of fit of the Paczynski curve to the baseline, where the baseline is defined as that part of the lightcurve that is more than twice the t_{FWHM} away from t_{MAX} . For this part of the lightcurve we use a cut of $\chi^2_{BL}/N < 1.5$, where N is the number of points in the baseline. The fits to the “empty” lightcurves that were used to assess the quality of the difference images showed that such a rather lenient cut is necessary, because our current photometry turns out to be quite sensitive to nearby variable objects which can cause some additional variability in some lightcurves.

• *Goodness of fit*

Finally, the shape of the peak must be consistent with microlensing, meaning that the Paczynski function must give a good fit. For the Monte Carlo simulations we find that a χ^2/N cut of 1.2 includes 90% of all events. However, secondary effects, like for example parallax effects, can influence the exact shape of the lightcurve. Also, the possibility of additional variability due to nearby variable objects must be taken into account. We use a rather lenient χ^2/N cut again, and since secondary effects will be stronger in high signal-to-noise events, this cut is also dependent on the “peak weight” P of the peak. The cut is described by $\chi^2_{PK}/N_{PK} < 1.5 + (\frac{P}{N_{PK}} - 1) \times 0.1$, where N_{PK} is the number of points in the peak, as defined above. Formally, a χ^2/N cut of 1.5 corresponds to a probability of 0.1% for the degrees-of-freedom in our fits, meaning that 99.9% of perfect microlensing lightcurves without any additional variability would pass this criterion.

A.2. Colour information

For several reasons the i' band data are important for the microlensing candidate selection. First, gravitational lensing is achromatic, so that the colour of the observed light in the difference images will not change during the microlensing event. Most periodic variable stars on the other hand, change colour during their pulsation cycles. Second, the long period variable stars like Mira’s that can otherwise easily be mistaken for microlensing events, are very cool and much brighter in i' than in r' . Also the variability is much larger and therefore easier to see in i' than in r' . Often, baselines that look flat in the r' data, clearly show variability in i' . Third, most long period variables turn out to have similar colours, providing an easy way to filter them out of the lightcurve database.

Paczynski fits are done to the two-season i' lightcurves of the microlensing candidates selected from the r' data. The values for the wavelength-independent lightcurve parameters t_{MAX} , t_{FWHM} and u_{MIN} are taken from the fits to the r' data and used in the i' fits, since the r' data constrain these parameters better because of their superior quality. In the case of large colour changes during the event, these Paczynski fits to the i' lightcurves should be poor, indicative of a non-microlensing event.

• *Colour cut*

As mentioned before, long period variable stars have cool atmospheres and therefore are very red. In Fig. A.2, the colours of the selected transients are shown as a function of the t_{FWHM}

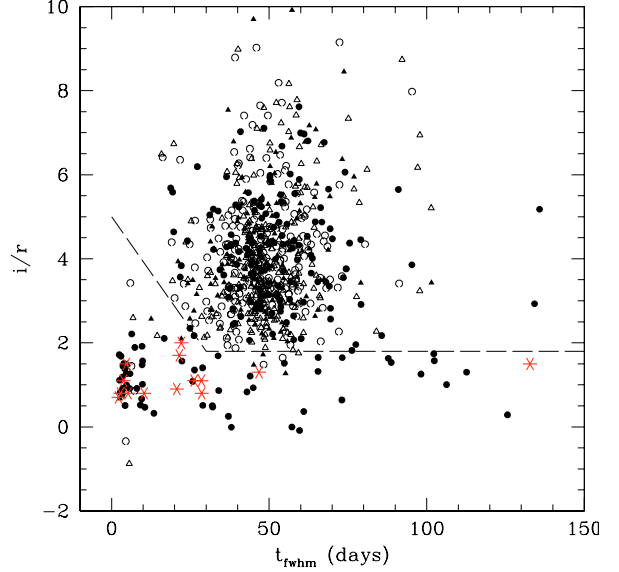


Fig. A.2. Plot of the peak width and colour of the peaks that pass our r' filtering procedure. The different symbols refer to the χ^2/N value of the Paczynski fits to the i' lightcurves. Closed circles are $\chi^2/N < 2$, open circles $2 < \chi^2/N < 3.5$, closed triangles $3.5 < \chi^2/N < 5$, and open triangles are $\chi^2/N > 5$. The positions of the 14 candidate microlensing events are indicated with stars. Clearly visible is that the long period variable stars cluster around a t_{FWHM} of 50 days and a i'/r' flux of 4. For further filtering, the colour cut indicated by the dashed line was used. A χ^2/N of 2 was used as the base value of the χ^2/N cut applied to the Paczynski fits to the i' data. The remaining sample still contains lightcurves with variable baselines and events caused by bad pixels.

of the Paczynski fits. The long period variable stars clearly stand out and cluster around a t_{FWHM} of 50 and an i'/r' flux ratio of 4. To remove the long period variables from the sample, a colour/ t_{FWHM} cut is used, indicated by the dashed line.

• *Goodness of fit*

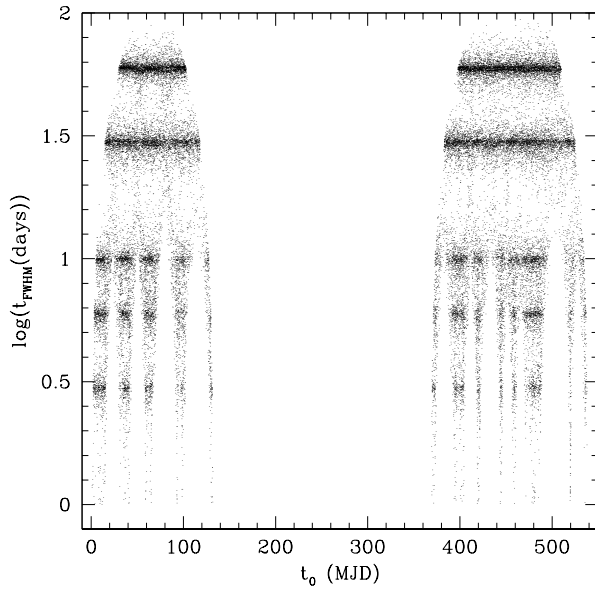
In the i' data, the problems of crowding of the residuals in the difference images is worse than in the r' data, meaning that our photometry is influenced by nearby variables even more than in the r' data. Since the long period variables are very red, more are detected in the i' difference images and the residuals are much brighter. In Fig. A.2 the transients are divided into four different bins depending on the χ^2/N value of the Paczynski fit to the i' data. A similar “peak weight” dependent χ^2/N cut was applied to the i' lightcurves as to the peak of the r' lightcurve, but with a minimum χ^2/N of 2 in stead of 1.5: $\chi^2/N < 2.0 + (\frac{P}{N_{PK}} - 1) \times 0.1$.

Appendix B: Monte Carlo simulations

In order to develop and test the efficiencies of our lightcurve filtering procedures to select microlensing lightcurves from a large set of lightcurves, we performed Monte Carlo simulations of microlensing events. To be meaningful for comparison with the data, the simulated lightcurves should have the same time sampling and error distribution as the data. Furthermore, a broad range of time scales, times of maximum amplification, and peak brightnesses has to be sampled.

Table B.1. The efficiencies of our r' lightcurve filtering procedure, in percent, for the simulated microlensing lightcurves.

t_{FWHM} (days)	Peak flux (ADU/s)						
	2	3.5	5	7	10	15	50
60	47	86	93	95	97	97	98
30	19	71	87	95	98	99	99
10	4	22	38	58	68	74	82
6	1	10	24	35	46	54	61
3	0	1	9	20	27	32	40
1	0	0	0	0	2	4	5

**Fig. B.1.** For all simulated microlensing lightcurves that pass our filtering procedure the fitted t_{FWHM} and t_{MAX} are plotted. From the clustering of the points it is clear that the fit program manages to recover the input widths of the simulations quite well. The gaps in the lightcurves show up as the events get shorter, while the period during which longer events can be detected is shorter than for the shorter events.

In total 170 000 lightcurves were simulated with random peak times, t_{FWHM} of 1, 3, 6, 10, 30 or 60 days, seven different peak fluxes, and impact parameters of 0.1, 0.3, 1.0 or 10. With these four impact parameters most of all possible lightcurve shapes are sampled, as is illustrated in Fig. A.1. For each simulated microlensing event, the time sampling and flux error bars were taken from a random real r' lightcurve, based on the 99/00 and 00/01 data. In this way, each simulated event has realistic characteristics, and the total sample of simulated events has the same observing characteristics as the total set of observed r' lightcurves. Standard Paczynski microlensing lightcurves were constructed, using the Gaussian error bars taken from the real data lightcurve.

In Table B.1 the efficiencies of the selection of microlensing lightcurves are tabulated for all combinations of t_{FWHM} and peak flux of the simulated events. For these detectability calculations simulated microlensing events were used for which the part of the peak higher than half of the peak flux lay completely within one observing season, which is one of our

Table B.2. The efficiencies of our r' lightcurve filtering procedure, in percent, for the simulated Mira lightcurves.

Period (days)	Amplitude (ADU/s)				
	5	10	15	20	30
150	5	2	1	0	0
200	5	3	2	2	1
250	34	26	24	20	17
300	25	22	19	17	15
350	2	1	1	1	0
400	0	0	0	0	0
450	6	4	3	2	2
500	17	12	10	8	8
550	29	23	21	17	20

selection criteria. It is clear that events with a peak flux lower than 3 ADU/s will hardly be selected. Also, events with a t_{FWHM} lower than 10 days are relatively problematic. This large decrease in detectability for short timescale events is primarily caused by gaps in the time sampling of our lightcurves. This is shown in Fig. B.1, where the fitted t_{FWHM} and t_{MAX} of the simulated microlensing lightcurves that pass the filtering procedure are plotted.

Since long period variable stars like Mira's can mimic microlensing lightcurves, it is important to determine to what extent these lightcurves are selected by our microlensing selection procedures. Although Miras do not have a flat baseline on a logarithmic plot, on a linear flux scale the peaks are much sharper and the minima very near zero, thus they are typically below our noise level except for short periods around the peak. Mira variables have periods ranging from 130 to 500 days, with the most typical period being 270 days (Petit 1987). However, there are a number of Mira's known with even longer periods (e.g. Jura et al. 1993; Rosino et al. 1997). Monte Carlo simulations of 120 000 Mira-type lightcurves were made with periods ranging from 150 to 550 days. The lightcurves were assumed to have a symmetric sawtooth shape in magnitude, with intrinsic variability amplitudes of 2, 4 or 6 mag and 5 different flux variability amplitudes. Time sampling and flux errors were taken from real data lightcurves in the same way as for the microlensing simulations and starting phase of the variability was chosen randomly.

The filter efficiencies for the selection of simulated Mira lightcurves are tabulated in Table B.2. Variable stars with periods up to 200 days are not likely to be selected and confused for microlensing events, because at least two peaks will always be present in the lightcurve. Periods between 200 and 350 days are clearly a problem for our 2 year dataset. This is caused by the time coverage of our lightcurves, which for these periods often gives a peak in one of the observing seasons and a second peak just in between the seasons. This is why using the data from the third season is important; variables with these periods that show only one peak in a two season lightcurve will show a second peak in the third season. Two examples of simulated Mira lightcurves with periods of 250 days that mimic microlensing and pass our filters are shown in Fig. B.2. Periodic variables with periods around 350 to 450 days are not selected because

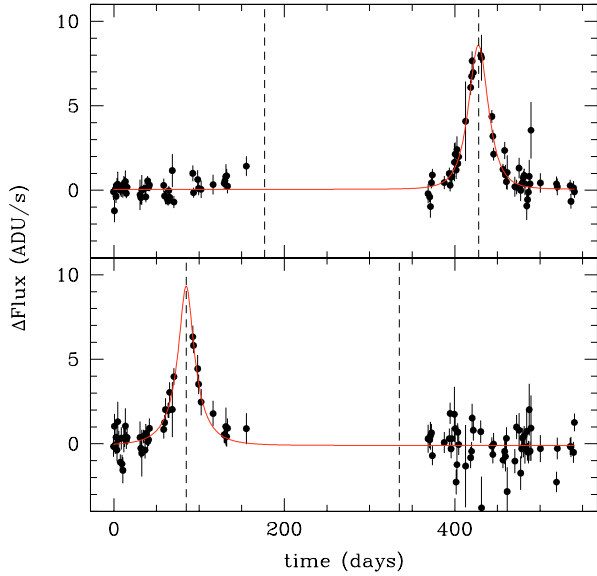


Fig. B.2. Two examples of simulated Mira lightcurves with periods of 250 days that pass our filters. The solid curves show the Paczynski fit to the lightcurves, and the dashed vertical lines indicate when the simulated Mira peaks. Both of the lightcurves have peaks in only one of the observing seasons, and therefore mimic microlensing quite well. A third observing season solves this problem for variables with periods around 250 to 300 days, but not very extremely long periods, i.e. longer than 500 days. Trends in the baseline like at the end of the first season in the upper lightcurve are often stronger in i' band than in r' band.

the fits are never very good and the baselines not flat. This is because a second peak will always have wings extending into one of the observing seasons. Longer period variables, however, can have only one peak in the lightcurve and a baseline that seems flat. Unfortunately, even with the third observing season data, these variables can still show one single peak and an otherwise flat baseline. Other characteristics, like colour and peak shape, have to be used to distinguish these variables from genuine microlensing events. For variable stars of all periods, the highest amplitude variables are least likely to be mistaken for microlensing events, because the difference in shape of the peak is detected better with higher amplitude.

The t_{FWHM} values of the simulated Miras show that variables with longer periods have broader peaks, as can be expected. However, even the longest timescale simulated Miras

show average t_{FWHM} values between 30 and 40, while the real variables cluster around a t_{FWHM} of 50 (Fig. A.2). Most detected real variables turn out to have more irregular peak shapes and lower peak heights than our simulated Miras, so that in reality long period variable stars are less likely to pass through our filtering procedure than Table B.2 suggests.

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